

1 **Reaction times can reflect habits rather than computations**

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22 **ABSTRACT**

23

24 Reaction times (RTs) are assumed to reflect the underlying computations required for making  
25 decisions and preparing actions. However, recent work has shown that movements can be  
26 initiated earlier than typically expressed without affecting performance; hence, the RT may be  
27 modulated by factors other than computation time. Consistent with that view, we demonstrated  
28 that RTs are influenced by prior experience: when a previously performed task required a  
29 specific RT to support task success, this biased the RTs in future tasks. This effect is similar to  
30 the use-dependent biases observed for other movement parameters such as speed or direction.  
31 Moreover, kinematic analyses revealed that these RT biases could occur without changing the  
32 underlying computations required to perform the action. Thus the RT is not solely determined by  
33 computational requirements but is an independent parameter that can be habitually set by prior  
34 experience.

35

36 **Impact Statement**

37 Although reaction times are often assumed to reflect the time required to select and prepare a  
38 movement, they can be strongly influenced by prior experience.

39

40 **Keywords**

41 reaction time; motor planning; movement initiation; use-dependent learning

42

## 43 INTRODUCTION

44

45         The reaction time (RT) is arguably the most widely used measure in neuroscience and  
46 psychology for noninvasively assessing processing in the brain: it is assumed to reflect the time  
47 needed to complete the perceptual and motor-planning computations required to prepare a  
48 response (Donders, 1969, Sternberg, 1969, Friston et al., 1996, Spivey, 2007, Sanders, 1998).  
49 This assumption appears justified by evidence that RTs are modulated by factors such as  
50 stimulus complexity (Kaswan and Young, 1965, Hick, 1952, Ratcliff, 2002), stimulus-response  
51 compatibility (Fitts and Seeger, 1953, Simon and Rudell, 1967, Simon and Wolf, 1963), number  
52 of potential responses (Henry and Rogers, 1960, Fischman, 1984, Christina et al., 1982), or  
53 required response accuracy (Fitts, 1954, Fitts, 1966, Reddi and Carpenter, 2000). As a  
54 consequence of this assumption, any reduction in the RT should negatively affect the quality of  
55 the resulting response. Indeed, instructing participants to respond as rapidly as possible leads to a  
56 decrease in accuracy (Fitts, 1966), consistent with changes in accuracy that occur through more  
57 direct manipulation of allowed preparation time (Schouten and Bekker, 1967, Ghez et al., 1997,  
58 Stanford et al., 2010, Haith et al., 2016).

59         However, recent evidence shows that it is possible to reduce the RT to a large extent  
60 before any decline in movement accuracy is observed. For example, when individuals are placed  
61 under strict time constraints, movement accuracy and task success decline only after the RT is  
62 shortened by ~80 ms (Haith et al., 2016). Similarly, startle by a loud acoustic stimulus can evoke  
63 initiation of a fully prepared action ~70 ms earlier than typically observed (Valls-Sole et al.,  
64 1999, Carlsen et al., 2004). These findings suggest that the RT includes some additional time that

65 is not required for computing the upcoming response, raising the possibility that other factors  
66 might also influence the RT.

67         An additional clue that the RT does not strictly represent computation time comes from  
68 evidence that the RT is influenced by context. This was illustrated, for example, in a recent study  
69 examining the planning of intentionally curved reaches (Wong et al., 2016). In this study, the  
70 RTs of simple point-to-point reaches were observed to be shorter than those of more complex  
71 curved movements, consistent with the idea that RT reflects computation time. However, when  
72 point-to-point movements were interleaved among curved reaches, the RTs of those point-to-  
73 point reaches were surprisingly prolonged by ~90 ms, matching the RTs of the curved reaches. A  
74 similar contextual effect has been observed during mental rotation: during a letter discrimination  
75 task, identifying letters that are presented upright occurs at much shorter RTs when all the letters  
76 are upright compared to when other letters in the same block of trials are rotated (Ilan and Miller,  
77 1994). These data suggest that RTs might be subject to experience-dependent biases – akin to  
78 other movement parameters such as speed or direction (Diedrichsen et al., 2010, Verstynen and  
79 Sabes, 2011, Hammerbeck et al., 2014, Huang et al., 2011) – rather than arising strictly from the  
80 outcome of computational requirements. That is, RTs may be subject to habit in the sense that an  
81 RT of a given magnitude may become more likely to be generated in the future simply because it  
82 has been generated in the past, regardless of current task requirements.

83         Here we present results from two experiments which demonstrate that the RT is subject  
84 to experience-dependent effects. In the first experiment – a target-interception task – we showed  
85 that experience with initiating reaches at a short or long RT exerts a corresponding bias on the  
86 RTs of subsequently performed point-to-point reaches.

87           In a second experiment, we demonstrated that these RT biases can also occur for more  
88 complex curved reaches around barriers. This task can be performed with or without the  
89 presence of a direct cue illustrating the appropriate trajectory, corresponding to greater or lower  
90 computational requirements respectively (Wong et al., 2016). We found that participants' RTs in  
91 the cued condition depended strongly on whether or not they had previously experienced the  
92 uncued condition. Analysis of the kinematics of these curved reaches allowed us to determine  
93 that the observed experience-dependent changes in RT were not due to habitually ignoring the  
94 cue and planning the movements in a different manner, but instead to habitually adopting a  
95 longer RT despite using a more rapid computation (i.e., taking advantage of the cue) to solve the  
96 task. Together, these two experiments reveal that the RT does not strictly reflect the time needed  
97 to complete the computations required for preparing responses, but may instead be selected  
98 habitually according to prior experience.

99

## 100 **RESULTS**

101

### 102 **Experiment 1: RTs were biased to be longer or shorter according to performance in a** 103 **previous task**

104

105           In Experiment 1, we tested whether performance of a task that encouraged movements to  
106 be generated with particular RTs could affect the RTs of subsequent movements performed in a  
107 different context. We asked participants to perform an interception task to hit a target that moved  
108 in a straight line toward or away from the participant (Fig. 1A). For one group of 10 participants,  
109 the target always moved outward, encouraging participants to initiate reaches at shorter RTs to

110 intercept the target before it moved beyond an invisible boundary and disappeared (see  
111 Methods). For a second group of 10 participants, the target always moved inward. This  
112 encouraged participants to increase their RTs, since they could reduce the effort required to  
113 complete this task by waiting for the target to move closer before reaching out to hit it. In both  
114 cases, participants were given no specific RT instruction. Additionally, because participants were  
115 not required to stop inside the target on interception trials, participants typically generated  
116 shooting movements through the target. Immediately before and after participants performed this  
117 interception task, we measured their RTs for simple point-to-point reaches (wherein the hand had  
118 to stop inside the target and move at a constrained speed) to assess any experience-related  
119 changes in RT.

120         At baseline, participants in both groups generated point-to-point reaches with comparable  
121 RTs (Table 1; no significant difference between groups:  $t = -0.22$ ,  $p = 0.82$ ), and were able to  
122 satisfy the requested speed requirements (see Table 1 and Methods).

123         Movement of the target, either toward or away from the participant's starting position,  
124 exerted a reliable influence on behavior during performance of the interception task. When the  
125 target moved outward (Fig. 1B; Figure 1 source data – Outward; Table 1), participants increased  
126 their movement speed (paired t-test comparing pre-training to the last training block,  $t(9) = 6.30$ ,  
127  $p < 0.001$ ) and decreased their RT (paired t-test comparing pre-training to the last training block,  
128  $t(9) = 2.99$ ,  $p = 0.02$ ) relative to their behavior on baseline point-to-point reaches. These  
129 reductions in RT persisted in the final phase of the experiment when the target was again  
130 stationary: RTs during the post-training point-to-point reaching block were shorter than in the  
131 analogous pre-training block by  $21.68 \pm 6.32$  ms (Fig. 1C; paired t-test comparing post to pre,  
132  $t(9) = -3.43$ ,  $p = 0.02$ ). This change in RT from pre-training to post-training was on average

133 76.6% of the total reduction in RT associated with the outward-interception task (i.e., pre-  
134 training compared to the last outward-interception training block). Importantly, these changes in  
135 RT for point-to-point movements occurred with no significant change in either peak velocity  
136 (paired t-test comparing post to pre,  $t(9) = 0.98, p = 0.70$ ) or endpoint error (paired t-test  
137 comparing post to pre,  $t(9) = 0.95, p = 0.70$ ). Thus, repeatedly generating movements at low RTs  
138 led to a reduction of the RT on subsequent point-to-point reaches with no detectable decrement  
139 in performance.

140 Interception of an inward-moving target exerted the opposite effect on behavior (Fig  
141 1D,E; Figure 1 source data – Inward; Table 1). Relative to baseline point-to-point movements,  
142 training on the inward-interception task led to a decrease of movement speed (paired t-test  
143 comparing pre-training to the last training block,  $t(9) = -6.82, p < 0.001$ ) and an increase in RT  
144 (paired t-test comparing pre-training to the last training block,  $t(9) = 2.45, p = 0.04$ ). RTs during  
145 a subsequent block of point-to-point reaches increased on average by  $12.87 \pm 4.94$  ms relative to  
146 baseline, although this change was not significant (paired t-test comparing post to pre,  $t(9) =$   
147  $2.60, p = 0.09$ ) and corresponded to only 9.69% of the increase in RT observed during the  
148 interception task. There was also a trend for the velocity of these reaches to increase although  
149 this also was not significant (paired t-test comparing post to pre,  $t(9) = 2.07, p = 0.14$ ) and there  
150 was no clear change in endpoint error (paired t-test comparing post to pre,  $t(9) = -0.56, p = 0.59$ ).

151 A mixed-effects model comparing the change in point-to-point RTs across the two groups  
152 revealed a significant interaction (likelihood ratio test;  $\chi^2(1) = 77.55, p < 0.001$ ) between the  
153 direction of target motion during training (outward or inward) and the change in RT from pre-  
154 training to post-training blocks (Fig. 1F). This interaction was driven by a difference in RT only  
155 during the post-training block (post-hoc test,  $p = 0.03$ ), since RTs during the pre-training block

156 were comparable. Together, these data reveal that the RT can be biased in an experience-  
157 dependent manner, suggesting that the RT does not strictly reflect obligate computation time.

158

## 159 **Experiment 2: RT biases occurred despite task-dependent changes in movement planning**

160

161 To further explore the experience-dependent bias in RT, we examined this effect in a  
162 more complex task that required participants to plan curved reaches around barriers. This task  
163 can be performed in one of two ways: either in the absence or presence of cues that illustrate a  
164 specific path around the barriers (Fig. 2A). Although the movements required to complete the  
165 task are identical in both cases, the presence of a path cue has been observed to provide a  
166 significant RT advantage (Wong et al., 2016). This advantage is thought to reflect a difference in  
167 computational requirements between the cued and uncued conditions, with the latter condition  
168 requiring an additional trajectory-planning stage to represent the path shape to be executed.  
169 Hence, this task provides an assay of habitual effects on RT between two sets of movements that  
170 have different planning requirements but the same execution requirements.

171

### 172 Experience-dependent RT biases were observed for curved reaches around barriers

173

174 In Experiment 2A, two groups of 12 participants were asked to generate reaches around  
175 barriers in the presence or absence of path cues. Comparable to previous findings (Wong et al.,  
176 2016), we observed that participants provided with a path cue exhibited a large RT advantage of  
177 52.94 ms compared to individuals performing the task without the path cue (initial cueing  
178 condition experienced by each group; uncued:  $349.76 \pm 8.81$  ms; cued:  $296.82 \pm 7.68$  ms; main

179 effect of condition,  $\chi^2(1) = 500.61, p < 0.001$ ; Fig. 2C; Figure 2 source data). Although the RT  
180 difference is smaller than observed previously (Wong et al., 2016), this is likely attributable to  
181 the fact that there were fewer possible barrier configurations presented in the current task  
182 compared to the previous version, i.e. consistent with Hick's law (Hick, 1952).

183 Each group was then exposed to a block of trials with the opposite cue condition (Fig.  
184 2B); that is, participants who first performed cued trials were next asked to perform the task  
185 without path cues, and vice versa. This switch revealed a similar experience-dependent effect as  
186 in Experiment 1. Specifically, whereas there was an RT advantage associated with the  
187 availability of a cue during the initial training block, when the conditions were switched the two  
188 groups exhibited similar RTs regardless of the presence or absence of the cue (cueing condition  
189 during the second block: uncued:  $334.75 \pm 8.17$  ms, cued:  $336.74 \pm 8.01$ ;  $\chi^2(1) = 0.035, p = 0.85$ ;  
190 Fig. 2C). This experience-dependent effect of path cues on the RT was confirmed by a  
191 significant main effect of the order in which the cued and uncued conditions were experienced  
192 ( $\chi^2(1) = 11.55, p < 0.001$ ), as well as a significant interaction between condition and order ( $\chi^2(1)$   
193  $= 105.42, p < 0.001$ ). This effect could not be attributed to changes in the RTs of uncued trials  
194 performed first compared to second ( $p = 0.25$ ); instead, this interaction arose because RTs were  
195 significantly longer on cued trials when participants had previously performed a block of uncued  
196 trials ( $p < 0.001$ ). Hence participants actually exhibited longer RTs on less computationally  
197 demanding cued trials if they experienced this condition second, even though they had far more  
198 experience performing a more challenging version of the task.

199 Interestingly, in Experiment 2A the RTs in the cued blocks (but not the uncued blocks)  
200 exhibited a slow asymptotic decline across trials (Fig. 2D). The rate of decline was not different  
201 between the two groups (non-significant interaction between group and trial,  $t = -1.46, p = 0.14$ ).

202 Power-law fits to the time series of latencies yielded similar exponents for both groups (cued  
203 trials experienced first: -0.025; cued trials experienced second: -0.029). This decline suggests  
204 that both groups were responding to the appearance of the path cues in a similar manner, perhaps  
205 by learning to take advantage of the cues when planning their movements (see the section  
206 below). Nevertheless, a constant RT offset remained across the entire block, reflecting a strong,  
207 persistent RT bias for the group that had prior experience with the uncued condition ( $t = 8.56$ ,  $p$   
208  $< 0.001$ ). Hence, even in this more complex task of generating reaches around barriers, RTs  
209 could be biased by prior experience.

210 In contrast to RT, task success in hitting the target while avoiding the barriers did not  
211 significantly vary with cueing condition or condition order (no effect of condition,  $\chi^2(1) = 2.26$ ,  $p$   
212  $= 0.13$ ; no effect of order,  $\chi^2(1) = 2.50$ ,  $p = 0.11$ ; non-significant interaction,  $\chi^2(1) = 1.63$ ,  $p =$   
213  $0.20$ ). That is, participants who were exposed to the uncued condition first did not become much  
214 more successful when they switched to the simpler, cued version of the task (uncued block:  
215  $81.97\% \pm 2.56\%$  successful; cued block:  $85.81\% \pm 1.62\%$  successful; posthoc test,  $p = 0.14$ ); nor  
216 did participants who were first exposed to the cued condition become less successful when  
217 switched to the more difficult uncued task (cued block:  $88.41\% \pm 1.66\%$  successful; uncued block:  
218  $88.02\% \pm 2.73\%$  successful; posthoc test,  $p = 1.00$ ). Note that in all cases participants could have  
219 further improved their task success. Thus while participants who initially performed the uncued  
220 condition exhibited prolonged RTs in the subsequent cued condition, this RT increase did not  
221 confer any performance advantage (as measured by task success) as might be expected by a  
222 speed-accuracy trade-off (Fitts, 1966).

223 We observed that only the RTs of cued trials were biased by having previously performed  
224 uncued trials, but no such bias was observed when performing these conditions in the opposite

225 order. Hence we wondered to what extent prior history influences this bias. Specifically, we  
226 asked whether participants who had prior experience performing cued trials at low RT would still  
227 exhibit an increase in cued RTs after completing an intervening block of uncued trials, or would  
228 they instead always express low RTs on cued trials once they had learned to perform them at low  
229 RTs. A separate group of 12 participants (Experiment 2B; Fig. 3; Figure 3 source data) were  
230 therefore asked to complete three blocks of trials: cued, uncued, and a second block of cued.  
231 Note, the cued blocks in this experiment have an analogous role to the pre-training and post-  
232 training assessment blocks of point-to-point reaches in Experiment 1.

233 Participants in Experiment 2B initially generated cued reaches at low RTs (pre-exposure  
234 cued block:  $349.81 \pm 36.03$  ms); however, experience with performing a block of uncued trials  
235 (uncued-block RT:  $390.26 \pm 66.37$  ms) led to prolonged RTs in the second cued block (post-  
236 exposure cued block:  $375.09 \pm 41.38$  ms; effect of block,  $\chi^2(2) = 532.46$ ,  $p < 0.001$ ; significant  
237 differences between all pairs of blocks,  $p < 0.001$ ; Fig. 3). This increase in RT from the pre-  
238 exposure to the post-exposure cued blocks reflected an average RT bias of 80.98% of the change  
239 in the RT from the initial cued to the uncued block (average change could not be calculated for  
240 one participant because that individual exhibited no change between the initial cued block and  
241 the uncued block). Thus, RT is strongly biased by prior exposure to the uncued condition in this  
242 task, even if participants had previously completed the cued task at low RT.

243

244 Path cues consistently influenced how movements were planned, independent of prior experience

245

246 Returning to Experiment 2A, we next asked what was the source of the RT bias. One  
247 possibility is that participants might simply have adopted a habitually long RT regardless of the

248 underlying computational demands of the current trial. Alternatively, participants that had  
249 initially been exposed to uncued trials might have ignored the cues when they became available  
250 and instead continued to plan the movement trajectories *de novo* – a more computationally  
251 demanding means of planning that requires prolonged RTs (Wong et al., 2016). To distinguish  
252 between these possibilities, we examined the kinematics of movements made with and without  
253 path cues, to find evidence of qualitatively different modes of movement planning. Specifically,  
254 we reasoned that the trajectories of movements planned using the cue should more closely  
255 resemble the path cues than movements planned *de novo*, regardless of the RTs at which these  
256 movements were initiated.

257         We found that the presence of path cues significantly influenced movement kinematics in  
258 a consistent manner across participants: the shapes of the average trajectories for the cued and  
259 uncued conditions (pairs of blue and green lines in Fig. 4A) were observed to be significantly  
260 different for at least one point along the length of the trajectory in 70% of all the target-barrier  
261 configurations performed (Fig. 4A, significant differences between trajectory pairs along the  
262 movement are denoted by yellow dots). For those trajectories where differences were observed,  
263 on average 21% of the movement was found to be significantly different depending on the  
264 condition. These differences often occurred near the start of the movement, wherein the hand  
265 passed closer to the starting barrier on cued trials. However, there was no significant effect of the  
266 order in which the conditions were performed; moreover, no kinematic differences were noted  
267 when comparing reaches performed in the same cueing condition (i.e., between cued trials  
268 experienced first or second, or between uncued trials experienced first or second). Thus cued  
269 reaches were always executed in a similar fashion regardless of whether participants had prior  
270 experience with the uncued condition.

271 In fact, a Procrustes shape-comparison analysis (Goodall, 1991; Fig. 4B; Figure 4 source  
272 data – Procrustes Distance) revealed that reaches significantly resembled the path cue more  
273 closely when it was available (effect of condition: -0.13, 95% confidence interval: [-0.19, -0.08],  
274 Bayes factor = 0). In contrast, there was no significant effect of order on trajectory shape (effect  
275 of order: -0.04 [-0.1, 0.02], Bayes factor = 14.24), even when comparing only cued trials (effect  
276 of order: 0.04 [-0.02, 0.09], Bayes factor = 13.63). Together, these data suggest that participants  
277 used the path cue to simplify planning their movements whenever it was available (Wong et al.,  
278 2016); that is, participants were changing how they planned their movements in response to the  
279 cue, which resulted in quantitative changes in the kinematics of their actions such that  
280 movements more closely resembled the cued paths. Similar effects were observed in Experiment  
281 2B (Source data for Experiment 2B kinematics analysis); in fact, the kinematics in the second  
282 cued block resembled the path cues more closely than even those of the initial cued block (effect  
283 of block: -0.04, 95% confidence intervals [-0.07, -0.02], Bayes factor = 0.39) – a finding that  
284 would be expected to reduce RTs as participants took advantage of the cue, not prolong them  
285 (e.g., consistent with Experiment 2A). Hence despite evidence suggesting that reaches were  
286 planned in an experience-independent manner according to whether a cue was present only on  
287 the current trial (in particular by relying on the cue to plan the action), RTs did not consistently  
288 reflect the computational benefit afforded by the path cue. Instead, RTs exhibited an experience-  
289 dependent bias that was unrelated to how these reaches were being planned and executed.

290

## 291 **DISCUSSION**

292

293           Although the RT is typically assumed to directly reflect the computation time required to  
294 plan a movement (Donders, 1969, Sternberg, 1969, Friston et al., 1996, Spivey, 2007, Sanders,  
295 1998), we demonstrated here that the RT may instead be influenced by previous experience. This  
296 was not simply a practice effect leading to more efficient planning of the same action; RT was  
297 influenced by recently generated RTs in the past even when those RTs were produced during  
298 another task with a different type of movement (e.g., in Experiment 1 the RTs of shooting  
299 movements to intercept a target biased the RTs of subsequent point-to-point reaches to stationary  
300 targets). This effect, while small, is comparable in magnitude to the RT changes previously  
301 observed in response to the addition of a delay period between the stimulus and go cues  
302 (Churchland et al., 2006), or to changes in stimulus or response complexity (Fitts, 1954, Fitts,  
303 1966, Hick, 1952, Simon, 1967, Henry and Rogers, 1960, for review, see Teichner and Krebs,  
304 1974). Moreover, we showed that RTs can be modulated to be either shorter or longer than  
305 previously expressed depending on the nature of the interposed task, and that these effects  
306 persisted across a large number of subsequent trials. Thus in general, the RT expressed on any  
307 given trial may be lengthened or shortened (by 20-30 ms) strictly due to prior experience, with  
308 no measurable impact on performance.

309

### 310 **Changes in motivation cannot explain RT biases**

311

312           Experiment 1 provides clear evidence that the RT can be modulated in an experience-  
313 dependent manner as if by a habit, and that such biases persist over a long period of time.  
314 However, this paradigm was unable to ascertain whether these RT biases arose because of  
315 changes in the rate of computational processing, or because of manipulations on the non-

316 computational portion of the RT. In the former case, changing the duration of computational  
317 processes could increase task success, particularly when shortening the RT in the outward-  
318 interception task. Thus there was a motivational incentive that could drive changes in the speed-  
319 accuracy trade-off in favor of shortened RTs while maintaining consistent accuracy (e.g.,  
320 motivation could have driven an increase in the rate at which evidence accumulated in a drift-  
321 diffusion model of RT (Ratcliff, 1978)), much in the same way that reward can reduce RT  
322 without affecting accuracy (Takikawa et al., 2002, Hubner and Schlosser, 2010, Stanford et al.,  
323 2010, Manohar et al., 2015). This modified speed-accuracy trade-off may then have been  
324 retained when planning subsequent point-to-point movements. Alternatively, if RT biases simply  
325 reflect modulation of the non-computational portion of the RT, the manner in which movements  
326 are planned should remain unchanged.

327         Two pieces of evidence from Experiment 1 speak against the hypothesis that the  
328 observed changes in RT occurred through a motivational effect. First, the RT bias persisted  
329 across a large number of trials (e.g., the entire next block) with no obvious decay toward the  
330 initially expressed RT. This persistence stands in contrast to the short-lived effects of motivation,  
331 which appear to decay after only a few subsequent trials (Takikawa et al., 2002, Xu-Wilson et  
332 al., 2009, Wong et al., 2015). Second, the effect of motivation typically influences not just the  
333 RT, but also movement speed and accuracy; in general, these three parameters have often been  
334 observed to modulate together in response to motivation (Takikawa et al., 2002, Wong et al.,  
335 2015). However, we observed no obvious changes in other movement kinematics such as  
336 movement speed or accuracy that accompanied the RT biases. On the other hand, speed and  
337 accuracy were constrained by task requirements during the point-to-point movement blocks.

338 Nevertheless, these data suggest that motivation alone cannot account for the long-lasting RT  
339 biases observed in Experiment 1.

340

### 341 **Habitual initiation rather than habitual planning**

342

343 The kinematically more complex trajectories required in Experiment 2 provided us with  
344 greater sensitivity to examine whether RT biases arose from effects on the computational or the  
345 non-computational portion of the RT. That is, we were better able to detect any execution-related  
346 differences that may have arisen from subtle changes in movement planning that were associated  
347 with changes in RT.

348 During Experiment 2, participants in the cued condition could have initiated their  
349 movements at short RTs; however, they exhibited persistently prolonged RTs when previously  
350 exposed to the more computationally demanding uncued condition. Such a bias in RT could have  
351 occurred for one of two reasons. Participants may have persisted in planning trajectories in the  
352 same manner as they had done without path cues, despite the availability of the path cue that  
353 would eliminate the need for this planning stage (*habitual planning*). Alternatively, participants  
354 may have used the path cue to reduce the computational load for planning but simply did not  
355 express their prepared actions earlier; that is, perhaps the RT does not strictly reflect computation  
356 time, but instead contains a manipulable, non-computational component that may be subject to  
357 habit (*habitual initiation*).

358 The analysis of movement kinematics allowed us to distinguish between these two  
359 alternatives. In particular, we demonstrated that movement kinematics changed in a consistent  
360 manner in the presence or absence of the path cues respectively, arguing that differences in

361 movement kinematics do reflect changes in how these actions were being planned. However, we  
362 observed no effect of condition order on trajectory kinematics (i.e., reach kinematics were  
363 always influenced similarly by the path cues when they were available, regardless of whether  
364 participants had prior experience with uncued trials), suggesting that participants did not ever  
365 exhibit habitual cue-free preparation of the movement trajectory when generating reaches in the  
366 presence of a cue. This is particularly evident in Experiment 2B, when movement kinematics  
367 indicate a particularly strong reliance on the cue to plan the action in the second cued block even  
368 though the RT did not reflect such changes in movement planning. Moreover, reaches in the  
369 uncued and cued conditions were confined to different portions of the workspace, making it  
370 unlikely that participants habitually applied the identical motor plans from previously-performed  
371 uncued reaches. These findings stand in contrast to the RT, which did not consistently modulate  
372 according to the presence or absence of the path cue, but was instead experience-dependent.  
373 Hence these data suggest that unlike motor planning – which is determined by the current task at  
374 hand – movement onset may be influenced in a use-dependent manner consistent with the idea of  
375 habitual initiation, and therefore may not represent the actual time required to prepare a  
376 movement.

377         If the RT can be habitual, this implies that the time of movement initiation is not simply  
378 the end result of computational processing but may actually be represented as a separate  
379 movement parameter during the pre-movement period, analogous to movement speed. Indeed,  
380 previous work (Haith et al., 2016, Brown and Robbins, 1991) has suggested that movement  
381 initiation may be independent of movement planning. Such a distinction between planning and  
382 initiation is consistent with recent data suggesting that an initiation signature can be observed in  
383 motor cortex preceding movement onset in a manner that is independent of the specific action

384 being prepared (Kaufman et al., 2016). Our finding that movement initiation can be determined  
385 by prior experience, rather than the time when planning has completed, further supports this  
386 view.

387         Selection of RT based on prior experience rather than computation time implies that it  
388 should be possible to choose a RT that is too short to allow for all planning processes to be  
389 completed prior to movement initiation. Consistent with this idea, there indeed appear to be  
390 situations in which the RT is spontaneously chosen to be improperly short (Orban de Xivry et al.,  
391 2017, Haith et al., 2016). In such cases, however, online corrections and refinements that occur  
392 after movement initiation help ensure a successful movement outcome (Orban de Xivry et al.,  
393 2017, Kohen et al., 2017, Wong and Haith, 2017). Hence, in the event that the RT is biased to be  
394 shorter than the time required to complete all necessary decision-making and planning-related  
395 computations, the motor system has online correction mechanisms in place to maintain overall  
396 task success.

397

398 **A common feature of movement parameters is that they are subject to experience-**  
399 **dependent biases**

400

401         Experience-dependent biases have been previously demonstrated for movement  
402 parameters such as speed or direction. For example, repetition of movements at a particular  
403 speed or toward a particular direction of the workspace strongly influences the kinematics of  
404 subsequently performed reaches (Diedrichsen et al., 2010, Verstynen and Sabes, 2011,  
405 Hammerbeck et al., 2014, Huang et al., 2011). Repetition has also been shown to influence the  
406 direction of movement invoked by transcranial magnetic stimulation (Classen et al., 1998). In

407 these cases, however, the influence of prior experience on movement biases appears for the most  
408 part to be short-lived, unlike the long-lasting effects seen here in the RT.

409 Why experience-dependent biases occur is still unclear. It has previously been proposed  
410 that such biases may simply be a result of Hebbian learning (Butefisch et al., 2000, Butefisch et  
411 al., 2004); repetition of the same action presumably strengthens the neural circuits that give rise  
412 to that movement, making it more likely to be invoked in the future. From a theoretical  
413 standpoint, experience-dependent biases have been framed in the context of a normative  
414 Bayesian model in which repetition leads to the construction of a strong prior that influences the  
415 preparation of future responses (Verstynen and Sabes, 2011). Regardless of whether one uses a  
416 mechanistic or computational framework, however, explanations for experience-dependent  
417 biases rely on the assumption that the parameter in question – e.g., speed or direction – is  
418 represented as a movement parameter that can be specified prior to movement. Therefore,  
419 experience-dependent biases on the RT imply that the RT is not simply the passive consequence  
420 of the time required to complete computational processing, but may instead be represented as a  
421 separate movement parameter that can be subject to habit.

422

## 423 **Conclusions**

424

425 In summary, these data support the hypothesis that the RT should be considered a distinct  
426 parameter that is selected during the pre-movement period; that is, the RT (i.e., initiation) may in  
427 part reflect, but is not strictly dependent upon, underlying computational processes. Thus,  
428 caution must be taken when interpreting differences in RT as indicative of the existence of  
429 computational stages with smaller or larger processing demands. Such RT differences may

430 simply reflect carryover from RTs recently exhibited in the past (i.e., habit), regardless of  
431 computational demands.

432

433 **MATERIALS AND METHODS**

434

435 Fifty-six right-handed, adult (average age, 23.27 years old; 26 males) neurologically  
436 healthy participants were recruited for this study. Twenty individuals participated in Experiment  
437 1 (10 individuals in each group); 24 individuals participated in Experiment 2A (12 individuals  
438 per group), and 12 individuals participated in Experiment 2B. Group sizes were chosen based on  
439 findings from previous studies (Wong et al., 2016). All participants provided written informed  
440 consent and were naive to the purposes of the study. Experimental methods were approved by  
441 the Johns Hopkins University School of Medicine institutional review board.

442 Participants made planar reaching movements with their right arm along the surface of a  
443 glass table. Their wrist was restrained in a wrist splint, and supported by pressurized air jets to  
444 allow frictionless movement of the elbow and shoulder. Vision of the arm was obscured by a  
445 mirror through which participants observed an LCD monitor (60 Hz), which displayed targets  
446 and a cursor representing the position of the index finger in a veridical horizontal plane.  
447 Movement of the index finger was tracked using a Flock of Birds magnetic tracking system  
448 (Ascension Technology, VT, USA) at 130 Hz.

449

450 **Experimental Paradigms**

451

452 Experiment 1: Interception Task

453

454 Targets appeared at any of 4 possible positions uniformly spaced every 90° along a circle  
455 from the start position. After a brief delay of 100 ms, the target immediately began moving, and

456 the participant was required to intercept the target before it vanished. In the outward-interception  
457 task, the target appeared 4 cm from the start position and moved outward until it was either  
458 intercepted or it reached 26 cm away from the start position, at which point it disappeared.  
459 Blocks consisted of 60 reaches (15 to each of the 4 target directions presented in a pseudo-  
460 randomized order); the target moved at a speed of 0.15 m/s in the first block, 0.225 m/s in the  
461 second block, and 0.3 m/s for the remaining 4 blocks. In the inward-interception task, the target  
462 appeared 24 cm away from the subject and began moving inward at a velocity of 0.12 m/s during  
463 each of the 5 blocks of 60 trials until it was intercepted or it reached the starting position.  
464 Participants had only 1 attempt to hit the target; the trial ended when the hand stopped moving  
465 (velocity < 0.05 m/s) or the hand exceeded the current target radius from the starting position.

466 Prior to and following these interception blocks, participants completed 60-trial blocks in  
467 which they were required to make point-to-point reaching movements to stationary targets  
468 located 15 cm away from the start position, uniformly spaced every 90°. Participants were also  
469 encouraged to satisfy a velocity criterion on each trial (between 0.6 and 0.9 m/s for the outward  
470 interception task, or between 0.5 and 0.8 m/s for the inward interception task), although no trials  
471 were excluded from analysis based on movement velocity. These differences in velocity criterion  
472 were deliberately chosen to encourage reaches at speeds that were comparable to those made  
473 during the interception training. Encouraged adherence to a velocity criterion simply allowed for  
474 an assay of changes in RT due to training on the interception task independently of any potential  
475 changes in movement speed. For both the pre- and post-training blocks, the first 10 trials were  
476 discarded from analysis to ensure that block averages for each participant reflected steady-state  
477 behavior.

478

479 Experiment 2: Barrier Task

480

481 Experimental methods were similar to a barrier-avoidance task reported previously  
482 (Wong et al., 2016) in which participants were required to generate a reaching movement toward  
483 a target while avoiding barriers that appeared around both the start position and the goal target.  
484 Targets appeared at any of eight positions uniformly spaced every 45° along a circle of radius 12  
485 cm; barriers consisted of three-sided boxes that could be oriented at any of eight angles (rotated  
486 in 45° increments). Participants initiated a trial by moving their hand into a central start target;  
487 after a random delay (600-1600 ms), the goal target and two barriers appeared (one around the  
488 start position and one around the goal target). Participants were instructed to begin their  
489 movement as soon as possible, and to reach toward the goal target while avoiding the barriers.  
490 As soon as participants initiated their reach, all visual information except the start target and the  
491 hand cursor disappeared. Participants stopped moving when they thought they reached the target  
492 location; upon conclusion of the movement they received visual feedback about their movement  
493 trajectory. In Experiment 2A, participants completed two blocks of 128 trials; in Experiment 2B,  
494 participants additionally completed a third block of 128 trials.

495 During one of the two blocks (or for Experiment 2B, the first and third blocks),  
496 participants were given a path cue in the form of a line drawn between the start and goal targets  
497 that navigated around the barriers. Additionally, in each block targets were confined to either the  
498 upper left or the lower right half of the screen such that all movements for a given block were  
499 confined to specific portions of the workspace. The direction of targets was different in the first  
500 and second block; thus, participants never performed the identical movements with and without a  
501 path cue, although cued and uncued reaches were rotationally symmetric (affording kinematic

502 comparisons). The target direction associated with a given block was counterbalanced across  
503 participants (in Experiment 2B the target direction was the same for the first and third blocks).

504 The presented path cues represented the average hand path taken to avoid the barriers  
505 from a prior experiment (Wong et al., 2016). The participants in Experiment 2A were subdivided  
506 evenly into two groups; in one group, participants performed uncued trials first to establish a  
507 history of performing a computationally difficult task, followed by cued blocks of trials (Group  
508 1: uncued trials first, cued trials second) while the remaining participants were given the opposite  
509 order of blocks (Group 2: cued trials first, uncued trials second). This allowed us to examine  
510 whether experience at performing reaches in the presence or absence of a path cue could bias  
511 future RTs. In Experiment 2B, all participants performed a block of cued trials, a block of  
512 uncued trials, and a final block of cued trials. Note that for both Experiment 2A and 2B, an equal  
513 number of participants within each group experienced cued reaches to the upper left as  
514 experienced cued reaches to the lower right; thus, kinematic differences across subjects could not  
515 be attributed to any movement direction-specific biases.

516

## 517 **Data Analysis**

518

519 Data were analyzed offline using programs written in MATLAB (The MathWorks,  
520 Natick, MA) and in R (R Core Team, 2016). Analysis code and data are available on GitHub at  
521 [https://github.com/BLAM-Lab-Projects/RT\\_habit](https://github.com/BLAM-Lab-Projects/RT_habit) (Wong, 2017). Reaches were selected  
522 according to a velocity criterion (tangential velocity greater than 0.05 m/s) and verified by visual  
523 inspection. For each movement, RT was computed as the time between target onset and  
524 movement initiation. Inherent delays in the system were estimated to be 105 ms on average; all

525 RTs have been corrected to compensate for this delay. Velocity was calculated by taking the  
526 numerical derivative of the hand position after smoothing using a second-order Savitzky-Golay  
527 filter with a frame size of 19 samples. Endpoint error was calculated as the absolute radial  
528 distance between the final position of the hand and the center of the target. All values are  
529 reported along with S.E.M.

530 In Experiment 1, RT, peak velocity, and interception amplitude were measured during  
531 interception training blocks. The amplitude of the target at the time of interception was  
532 calculated as the distance of the target away from the central starting position at the first time the  
533 hand entered the target; if the participant was not successful at intercepting the target, no target  
534 amplitude was recorded for that trial. During pre- and post-interception blocks, RT, peak  
535 velocity, and endpoint error were measured for point-to-point reaches. These three metrics were  
536 compared within groups using paired t-tests, with p-values adjusted for multiple comparisons  
537 using Bonferroni-Holm corrections. RT was also compared across groups using a mixed-effects  
538 model in R using the *lme4* package (Bates et al., 2015), with post-hoc pairwise tests performed  
539 using the generalized linear hypothesis testing function in the *multcomp* package (Hothorn et al.,  
540 2008) and adjusted for multiple comparisons using the Bonferroni-Holm correction.

541 In Experiment 2, trials were excluded if participants did not complete their reach within  
542 1200 ms. Additionally, since most barrier configurations had more than one possible solution  
543 (e.g. above or below a barrier), we pre-selected one possible path as “canonical” and presented  
544 that solution on path-cued trials; any “non-canonical” reaches were excluded to allow for a fair  
545 comparison of RTs for movements of comparable kinematics. Reaches were not excluded if  
546 participants simply hit one of the barriers or did not reach close enough to the target to be  
547 considered successful on that trial but otherwise satisfied the inclusion criteria noted above. On

548 average across participants, about 7.5% of reaches were excluded from uncued blocks and 3.0%  
549 of reaches were excluded from cued blocks.

550 RTs were compared in R with mixed-effects models using the *lme4* package (Bates et al.,  
551 2015). For Experiment 2A, this model treated group (uncued reaches first or cued reaches first)  
552 and condition (cued or uncued) as fixed effects and barrier configuration and participant as  
553 random effects. Significant effects were determined using a likelihood ratio test to compare pairs  
554 of models (with and without the factor of interest). In Experiment 2B, there was only a main  
555 effect of block (1, 2, or 3), with post hoc tests performed in R using the generalized linear  
556 hypothesis testing function in the *multcomp* package (Hothorn et al., 2008). All *p*-values  
557 obtained from post hoc tests were adjusted for multiple comparisons using Bonferroni-Holm  
558 corrections.

559 The time course of the change in RT during path-cued blocks was compared across  
560 groups by fitting generalized linear models in R with factors of group and trial using the *nlme*  
561 package (Pinheiro et al., 2014) to account for the autocorrelated covariance structure across time  
562 within participants. These models remove any autocorrelation structure across trials for each  
563 participant individually prior to examining main effects; the form of the autocorrelation structure  
564 was selected by fitting Autoregressive-Moving Average (ARMA) models to the data and  
565 selecting the model fit that yielded the lowest Aikake Information Criterion on average across  
566 participants; this led to a choice of an ARMA(1,1) process.

567 Movement kinematics were examined using two methods. First, movements were  
568 compared using tools of functional data analysis (Goldsmith and Kitago, 2015). Briefly,  
569 trajectories were time-normalized and evenly resampled. For each trajectory, mean pairwise  
570 differences were examined using a function-on-scalar regression model fit using a Bayesian

571 method that allows for correlations in the errors. A 95% simultaneous posterior credible interval  
572 was used to identify significant differences between conditions, accounting for the multiple  
573 comparisons made across time points within a single trajectory but not for multiple comparisons  
574 across trajectories.

575         Second, movement trajectories were also examined by comparing the shape similarity of  
576 each movement to the path cue using a Procrustes distance metric (Goodall, 1991). The  
577 Procrustes distance finds the best combination of translation, rotation, and scaling to match a  
578 shape to its template, then estimates the remaining dissimilarity normalized between 0 and 1,  
579 where 0 implies the two shapes are perfectly matched. For each movement, the Procrustes  
580 distance was estimated; then the overall distributions of Procrustes distances for the cued and  
581 uncued conditions were compared using a generalized linear mixed model with a log-normal link  
582 function using *brms*, an R interface to the Stan language (Buerkner, 2016, Carpenter et al., 2016,  
583 Hoffman and Gelman, 2014). This generalized linear mixed model had main effects of group  
584 (uncued reaches first or cued reaches first) and condition (cued or uncued); significant effects  
585 were estimated by calculating Bayes factors to test the null hypothesis that the effect coefficient  
586 is equal to zero (according to whether the confidence intervals of the effect included zero).  
587

588 **Figure Legends**

589

590 Figure 1. Experiment 1: interception task. (A) Paradigm. Participants were first asked to perform  
591 a block of point-to-point reaches. Following this, participants in the outward interception task  
592 were required to hit targets that moved away from them, while participants in the inward  
593 interception task were required to hit targets that moved toward them. Following this, all  
594 participants completed a final block of point-to-point reaches. (B) Time course of average RT,  
595 average peak velocity, and average amplitude changes during the outward-interception task.  
596 Peak velocities are compared to the actual target velocity (solid gray line) and reach amplitude is  
597 compared to the average target amplitude at interception (solid gray line; note for pre- and post-  
598 blocks the targets remained at a fixed amplitude). For reaches made during the interception task,  
599 the reported amplitude is the magnitude of the entire movement generated (i.e., until the velocity  
600 of the hand returned to zero) regardless of when or whether the target was successfully  
601 intercepted. Shaded regions represent S.E.M. (C) Comparison of baseline (blue) and post-  
602 training (green) performance for average RT, average peak velocity, and average endpoint error;  
603 each gray line is an individual participant. (D,E) As in panels A and B, but for the inward-  
604 interception task. (F) Across the two tasks, there was a significant interaction between task and  
605 block on average RT; this was driven by an RT shift away from baseline behavior in opposite  
606 directions for the two groups following training.

607

608 Figure 2. Experiment 2A: RTs depended on prior experience in the barrier task. (A) Participants  
609 reached from a start position to a target without intercepting a pair of barriers. One example  
610 target and barrier configuration is shown at left. Under the path-cued condition (at right),

611 participants were also provided with a cue (solid gray line) indicating how they should get to the  
612 target. Note, in both cases the barriers, target, and cue (when present) disappeared upon  
613 movement onset; hence participants in the cued condition were not merely tracing the cue but  
614 were required to hold the desired movement trajectory in memory, analogous to individuals  
615 performing the uncued condition. (B) Participants were divided into two groups; one group  
616 performed an uncued reaching block followed by a cued reaching block, while the other group  
617 performed these blocks in the opposite order. (C) Upper panel, the average RT across all  
618 participants for each condition (cued or uncued) for the two groups (uncued-first or cued-first).  
619 Lower panel, the average difference in RT between cued and uncued reaches across all barrier  
620 configurations; positive differences indicate that RTs were shorter on cued reaches compared to  
621 uncued reaches. Dots represent individual participants. (D) The time course of RTs across all  
622 trials, for each group (red, uncued-first group; orange; cued-first group). Black lines represent  
623 power-law fits to the decay of RT during the cued blocks of trials.

624

625 Figure 3. RTs in Experiment 2B. (A) Average RTs during each block, in the order that blocks  
626 were experienced in the experiment (from left to right). (B) Change in average RT between each  
627 pair of blocks during the experiment; each gray line is an individual participant.

628

629 Figure 4. Movement kinematics in Experiment 2A depend on the presence of the path cue. (A)  
630 Average reach trajectories across participants for each barrier configuration for uncued (blue)  
631 and cued (green) conditions. Each yellow dot represents a time at which the mean trajectories  
632 were found to be consistently different between conditions. Since no significant effect of order  
633 was observed, data were collapsed across groups. (B) Trajectory shapes for cued and uncued

634 reaches were examined for their similarity to the path cue according to a Procrustes distance  
635 analysis. A Procrustes distance of zero means the trajectory is identical in shape to the path cue.  
636 Distributions of estimated Procrustes distances pooled across all trajectory shapes were averaged  
637 across all participants in each group (uncued-first or cued-first) and are shown for comparison;  
638 the mode of each distribution is indicated by the vertical dashed line for visualization purposes.  
639 No significant effect of group was observed. (C) Average difference in Procrustes difference  
640 between uncued and cued reaches for each individual participant is shown; positive differences  
641 indicate that cued reaches are more similar in shape to the path cue compared to uncued reaches.  
642  
643

644 **Source Data Legends**

645

646 Figure 1 source data – Outward. This data set pertains to the outward-interception task, and were  
647 used to calculate the summary statistics in Figure 1C and Figure 1F. This file contains measured  
648 kinematic parameters (RT, peak velocity ( $V_{peak}$ ), and endpoint error (Err), for each trial and  
649 participant during the pre-training (Block 1) and post-training (Block 2) point-to-point reaching  
650 blocks. Note, the first 10 trials in each block have been excluded to evaluate steady-state  
651 movement parameters.

652

653 Figure 1 source data – Inward. This data set pertains to the inward-interception task, and were  
654 used to calculate the summary statistics in Figure 1E and Figure 1F. This file contains measured  
655 kinematic parameters (RT, peak velocity ( $V_{peak}$ ), and endpoint error (Err), for each trial and  
656 participant during the pre-training (Block 1) and post-training (Block 2) point-to-point reaching  
657 blocks. Note, the first 10 trials in each block have been excluded to evaluate steady-state  
658 movement parameters.

659

660 Figure 2 source data. This file contains RT data for Experiment 2A, used to generate Figure 2C  
661 and Figure 2D. These data were also used to fit mixed-effects models for computing statistics.  
662 The file contains the RT for each trial in each group (order, cued-first = 0, uncued-first = 1) and  
663 condition (cued = 0, uncued = 1). Additionally, there is a factor representing the barrier  
664 configuration (TB), which is assigned an arbitrary numbering from 0 to 63, which is used in the  
665 mixed-effects modeling analysis.

666

667

668 Figure 3 source data. This file contains RT data from Experiment 2B, used to generate Figure 3.

669 These data were also used to fit mixed-effects models for computing statistics. The file contains

670 the RT for each trial in every condition (cued block = T, uncued = NT), for all three blocks (first

671 cued block, uncued block, second cued block). Additionally, there is a factor representing the

672 barrier configuration (TB), which is assigned an arbitrary numbering from 0 to 63, which is used

673 in the mixed-effects modeling analysis.

674

675 Source data for Experiment 2B kinematics analysis. These data contain the Procrustes Distance

676 measure (range, 0 to 1) calculated for each trajectory performed by every participant in

677 Experiment 2B. The file includes a column for condition (cued block = 0, uncued = 1) and a

678 column for block (first cued block = 1, uncued block = 2, second cued block = 3). These data

679 were used to calculate kinematic differences between blocks using a generalized linear mixed

680 model in *brms* (see Methods). There is no corresponding figure for these data, summary statistics

681 are reported in the Results.

682

683 Figure 4 source data – Procrustes Distance. These data contain Procrustes distance estimates for

684 all the trajectories performed in Experiment 2A, used to generate Figure 4B and Figure 4C. The

685 file contains the Procrustes distance for every trial in each group (order, cued-first = 0, uncued-

686 first = 1) and condition (cued = 0, uncued = 1). These data were used to examine kinematic

687 differences between conditions and groups (i.e., an effect of order) using a generalized linear

688 mixed model in *brms* (see Methods).

689



691 Table 1: Movement parameters for Experiment 1

692

	Reaction time (ms)	Peak velocity (m/s)	Endpoint error (cm)
<b>Outward Interception</b>			
Pre-training	295.38 ± 14.98	0.66 ± 0.02	1.16 ± 0.10
Last training block	255.20 ± 7.85	1.28 ± 0.10	n/a
Post-training	273.20 ± 12.10	0.67 ± 0.01	1.32 ± 0.22
<b>Inward Interception</b>			
Pre-training	290.76 ± 18.35	0.55 ± 0.02	1.09 ± 0.06
Last training block	564.82 ± 112.11	0.27 ± 0.04	n/a
Post-training	303.63 ± 17.25	0.57 ± 0.01	1.05 ± 0.04

693

694 RT, velocity, and endpoint error for the two groups in Experiment 1 for point-to-point  
 695 movements measured before and after training, as well as for shooting movements during the last  
 696 block of training. Reach velocity was required to be between 0.6 and 0.9 m/s for the outward  
 697 interception task and between 0.5 and 0.8 m/s for the inward interception task. Note that no  
 698 endpoint error is reported for the last block of training because participants reached past a  
 699 continuously moving target, making it difficult to define error when participants missed the  
 700 target during those blocks.

701

702

703 **Conflict of Interest**

704 The authors declare no competing financial interests.

705

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709

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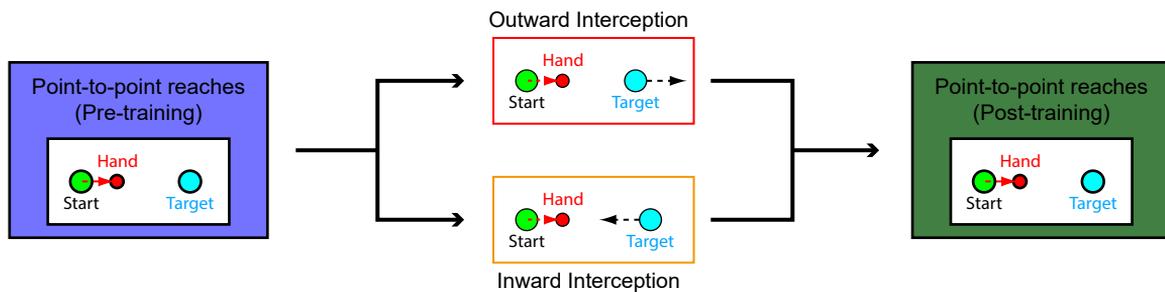
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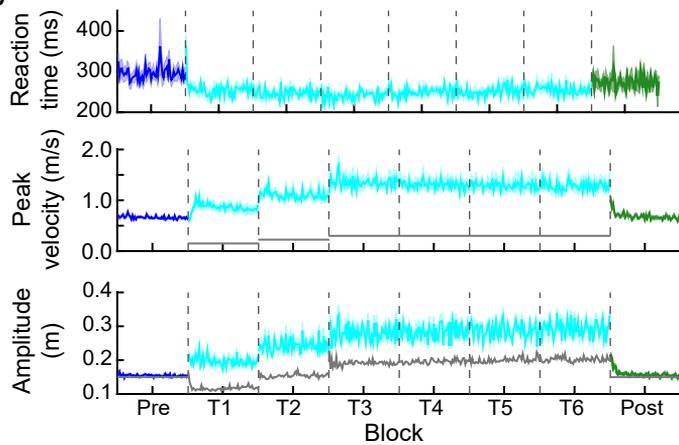
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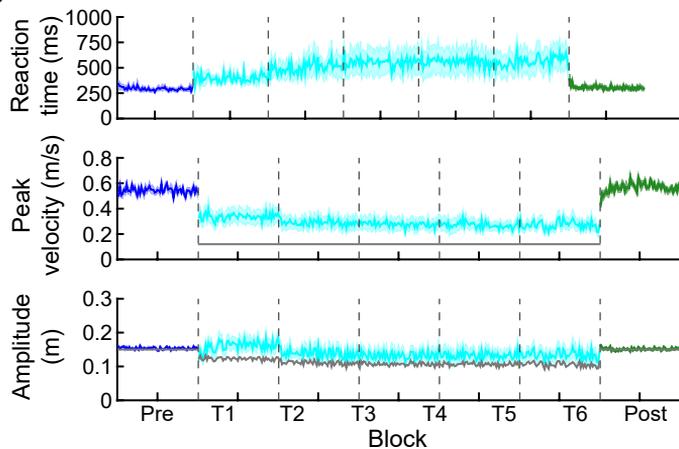
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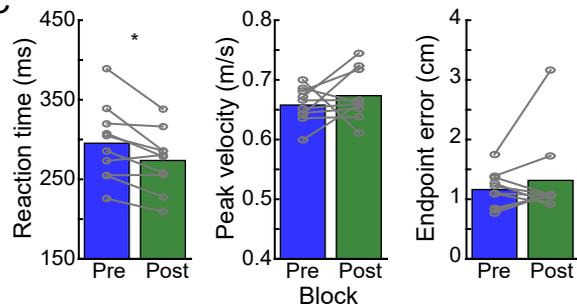
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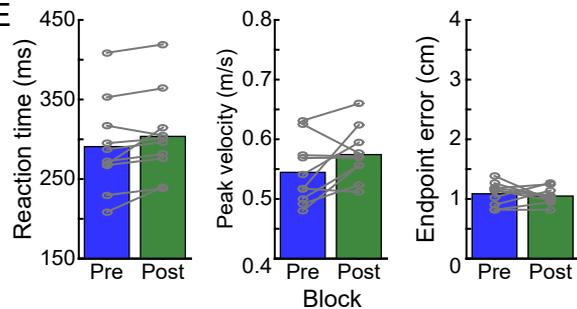
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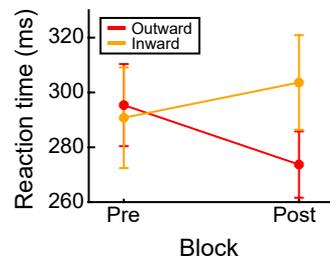
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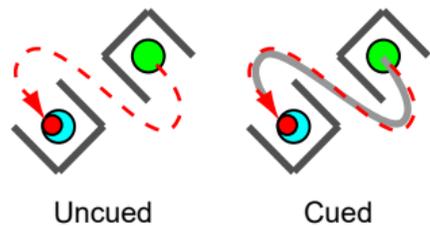
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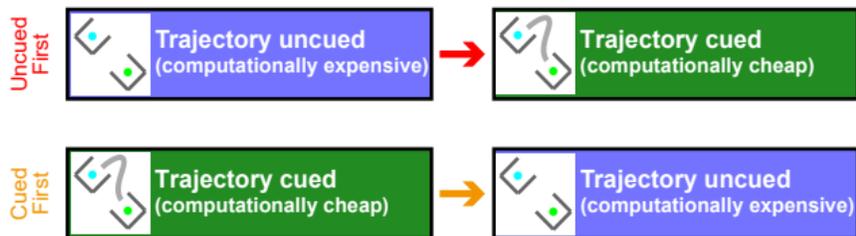
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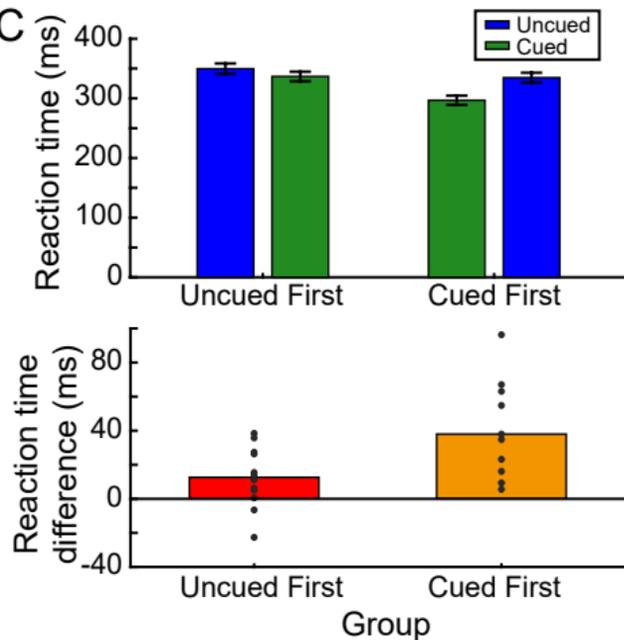
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